# Bootstrap and Linear regression

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STAT 3150—Statistical Computing

• Understand the difference between resampling cases vs residuals.

- In the last two modules, we reviewed linear regression and discussed residual analysis.
  - We discussed the linear regression assumptions, and their relative importance.
- In this module, we will discuss how to use bootstrap in the context of linear regression.
- There are actually 2 different ways of using bootstrap, corresponding to 2 different sets of assumptions concerning the data generating mechanism.

- When the error term is normally distributed, we know the distribution of the estimator  $\hat{\beta}$ :

$$\hat{\beta} \sim N\left(\beta, \sigma^2(\mathbb{X}^T\mathbb{X})^{-1}\right).$$

- This can be used to compute p-values and confidence intervals.
- But when we *don't know* the distribution, or if we *don't want* to assume it follows a normal distribution, we can use bootstrap to make valid inference.

### Bootstrap and Linear regression ii

- As we will see, there are two different ways to use bootstrap:
  - · Resample cases;
  - · Resample residuals.
- The main difference is how many assumptions we want to retain:
  - To resample residuals, we need to assume additivity, linearity, and homoscedasticity.
- In both cases, we still need to assume **independence of the errors**.

## **Resampling cases**

- This is the simplest form of bootstrap for linear regression.
  - It should also be familiar.
- For this form of bootstrap to be valid, we only need to assume the errors are independent.
- In fact, it can be shown that when resampling cases, the bootstrap estimate of the standard error is approximately equal to the Huber-White robust standard error.

### Algorithm (Cases)

- 1. Sample with replacement from  $(Y_1, \mathbf{X}_1), \ldots, (Y_n, \mathbf{X}_n)$ .
- 2. Refit the linear model using the bootstrap sample and obtain bootstrap estimates  $\hat{\beta}^{(b)}$ .

library(DAAG)
# Recall
fit1 <- lm(magnetic ~ chemical, data = ironslag)</pre>

```
str(boot_beta1)
```

```
## num [1:2, 1:5000] 7.1357 0.6105 -0.0984 1.0029
0.8562 ...
## - attr(*, "dimnames")=List of 2
## ..$ : chr [1:2] "(Intercept)" "chemical"
## ..$ : NULL
```

```
se_int <- sd(boot_beta1[1,])
se_slope <- sd(boot_beta1[2,])</pre>
```

```
cbind("Lower" = coef(fit1) - 1.96*c(se_int, se_slope),
        "Upper" = coef(fit1) + 1.96*c(se_int, se_slope))
```

## Lower Upper
## (Intercept) -3.2731976 6.078392
## chemical 0.6725151 1.159025

# Compare to MLE theory
confint(fit1)

## 2.5 % 97.5 %
## (Intercept) -3.7856893 6.590884
## chemical 0.6768355 1.154704

- Our confidence interval for the intercept is a bit smaller, but it still includes 0.
- On the other hand, the confidence interval for **chemical** is comparable to the one from MLE theory.

- As mentioned above, this approach requires more assumptions than resampling cases:
  - Additivity and linearity;
  - Homoscedasticity.
- But the trade-off is that we get smaller confidence intervals than if we resample cases.

### Algorithm (Residuals)

First, compute residuals  $E_i$  and fitted values  $\hat{Y}_i = \hat{\beta}^T \mathbf{X}_i$  for each observation i = 1, ..., n.

- 1. Sample with replacement from the residuals and obtain a bootstrap sample  $E_1^{(b)}, \ldots, E_n^{(b)}$ .
- 2. Add the bootstrapped residuals to the fitted values:  $Y_i^{(b)} = \hat{Y}_i + E_i^{(b)}.$
- 3. Using these new outcomes  $Y_i^{(b)}$  and the original covariates  $\mathbf{X}_i$ , fit a linear regression model and obtain bootstrap estimates  $\hat{\beta}^{(b)}$ .

```
library(MASS)
# Recall
dataset <- transform(mammals,
                          log_body = log(body),
                          log_brain = log(brain))
# Fit model
fit2 <- lm(log_brain ~ log_body, data = dataset)</pre>
```

```
# Compute residuals
resids <- resid(fit2)</pre>
n <- length(resids)</pre>
boot_beta2 <- replicate(5000, {</pre>
  indices <- sample(n, n, replace = TRUE)</pre>
  logbrain boot <- fitted(fit2) + resids[indices]</pre>
  fit boot <- lm(logbrain boot ~ log(mammals$body))</pre>
  coef(fit boot)
})
```

```
str(boot_beta2)
```

```
## num [1:2, 1:5000] 2.128 0.778 2.051 0.733
1.992 ...
## - attr(*, "dimnames")=List of 2
## ..$ : chr [1:2] "(Intercept)"
"log(mammals$body)"
## ..$ : NULL
```

```
se_int <- sd(boot_beta2[1,])
se_slope <- sd(boot_beta2[2,])</pre>
```

```
cbind("Lower" = coef(fit2) - 1.96*c(se_int, se_slope),
        "Upper" = coef(fit2) + 1.96*c(se_int, se_slope))
```

## Lower Upper ## (Intercept) 1.9517078 2.3178695 ## log\_body 0.6964323 0.8069395

```
# Compare to MLE theory
confint(fit2)
```

## 2.5 % 97.5 %
## (Intercept) 1.9426733 2.3269041
## log\_body 0.6947503 0.8086215

- This time, we can see that we get essentially the same result in both cases.
  - The bootstrap confidence intervals are slightly smaller.

- Note: Other types of residuals can be used for the bootstrap, e.g. to mitigate the effect of outliers.
  - But don't use standardized residuals! You want the residuals to retain approximately the same variance as in the original data.

- We looked at two different ways to perform bootstrap in the context of linear regression.
  - Resample the **cases** or the **residuals**.
- Resampling the cases is valid more generally than resampling the residuals.
- But resampling the residuals can lead to smaller, more accurate confidence intervals.
- Deciding which approach to use is a question of how much you trust the model.

- **Importantly**, neither approach is valid when the errors are *correlated*.
  - E.g. clustered data, repeated measurements, time series.
  - Bootstrap can be adapted to these methods, but this is beyond the scope of STAT 3150.